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On the Stochastic Approximation Method and Optimal Filtering Theory

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In this short paper we wish to establish some connections among the maximum likelihood estimate, the optimal filtering, and the stochastic approximation solutions to the following well-known problem: Consider the vector-matrix equations

$$Ax + v_k = b_k \quad k = 1, 2, \dots \quad (1)$$

where

A is a given $r \times n$ matrix

x is an unknown n -vector

v_k is a random r -vector with $E(v_k) = 0$

and

$$E(v_k v_j') = I\delta(k - j)$$

b_k is a r -vector of observation

One wishes to determine an estimate x for the unknown parameters x which is optimal in some sense.

First, we shall make a maximum likelihood estimate for x . It is well-known [1, 2] that a recursive method for calculating the estimate in the case of gaussian noise is

$$\begin{aligned} \hat{x}_{k+1} &= \hat{x}_k + P_{k+1}A'(b_{k+1} - A\hat{x}_k) \\ \hat{x}_0 &= 0 \end{aligned} \quad (2)$$

where

$$\begin{aligned} P_{k+1}^{-1} &= P_k^{-1} + A'A \\ P_0 &\text{ given } n \times n \text{ positive definite matrix} \end{aligned} \quad (3)$$

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Note that (3) only yields the inverse of P_{k+1} which is needed in (2). However, it is well known in numerical analysis, though not in control theory, that the following lemma is true [6]:

LEMMA. If $P_{k+1}^{-1} = P_k^{-1} + A'A$ where $P_k > 0$ then

$$P_{k+1} = P_k - P_k A' (A P_k A' + I)^{-1} A P_k \quad (4)$$

PROOF: by direct verification. Using (4) one can easily establish again by straightforward matrix manipulation that

$$P_{k+1} A' = P_k A' (A P_k A' + I)^{-1} \quad (5)$$

which leads to

$$\hat{x}_{k+1} = \hat{x}_k + P_k A' (A P_k A' + I)^{-1} (b_{k+1} - A \hat{x}_k) \quad (6)$$

Equations (4) and (6) are essentially the estimation scheme of Kalman based on the original Wiener's theory [3]. Bellman [4] has also derived (4) using dynamic programming.

Now repeated applications of (5) to itself lead to

$$\begin{aligned} P_k A' &= P_{k-1} A' (A P_{k-1} A' + I)^{-1} = P_{k-2} A' (2A P_{k-2} A' + I)^{-1} \\ &= P_0 A' (k A P_0 A' + I)^{-1} \end{aligned} \quad (7)$$

such that we have in the limit

$$\lim_{k \rightarrow \infty} P_k A' = \frac{1}{k} P_0 A' (A P_0 A')^{-1} \quad (8)$$

Thus, (6) becomes asymptotically

$$\hat{x}_{k+1} = \hat{x}_k + \frac{1}{k+1} P_0 A' (A P_0 A')^{-1} (b_{k+1} - A \hat{x}_k) \quad (9)$$

We recognize that (9) represents a form of multidimensional stochastic approximation [5] with the weighting function going down as $1/k$ as it is required. Furthermore, if we take $r = n$ (which represents no loss of generality inasmuch as we are investigating the case where $k \rightarrow \infty$) and assume that A^{-1} exist (which is the same as the usual observability condition [3]), then (9) becomes simply

$$\begin{aligned} \hat{x}_{k+1} &= \hat{x}_k + \frac{1}{k+1} [x - (\hat{x}_k + A^{-1} v_k)] \\ &= \frac{1}{k+1} \sum_{i=0}^{k+1} (x - A^{-1} v_i) = \frac{1}{k+1} \sum_{i=0}^{k+1} A^{-1} b_i \end{aligned} \quad (10)$$

which is just the weak law of large numbers.

Equations (4), (5), and (7) can be extended in many ways to include effect of time dependence, etc. The expressions become, however, more involved. For example, for $r = n$, (7) becomes

$$P_k A'_k = P_0 A'_1 \left[\left(\sum_{i=1}^k A_1^{-1} (A'_i A_i) \right) P_0 A'_1 + I \right]^{-1} A_1^{-1} A'_k \quad (11)$$

for the case where A is time varying. Note that the $1/k$ effect is nevertheless still present. The presence of P_0 in the formula allows us to take advantage of the confidence one has in the initial estimate \hat{x}_0 . While in the regular stochastic approximation scheme, this fact is missing and the weighting factor goes down as $1/k$ instead of $1/k + k_0$ where k_0 is a constant.

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